# Variability of Blood Constituents in University <br> Students: Correlations over a Three-Year Period, and Estimations of Annual Incidences and Continual Proportions of Abnormality 

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#### Abstract

Objectives. To estimate the annual abnormality incidence proportions and annual continual abnormality proportions of blood constituents in order to examine the variability of blood characteristics during young adulthood. Methods. We estimated the annual abnormality incidence proportions and annual continual abnormality proportions of blood constituents from 4th-year abnormality proportions of 1st-year abnormal subjects and from 4th-year abnormality proportions of 1st-year normal subjects. The estimation was performed only on items having 30 or more 1st-year abnormal subjects, using two calculation models. Results. Model I, which assumed the coming and going of abnormal subjects, was more suited for the estimation than Model 2, which ignored coming and going. The blood constituents of university students had the following characteristics: i) 4th-year abnormality proportions were higher than 1st-year abnormality proportions for many items; however, except for serum protein and NF, the differences were not significant; ii) 4th-year abnormality proportions of 1st-year abnormal subjects, compared to 4th-year abnormality proportions of 1st-year normal subjects, were twice as high for low items and over 10 times as high for high items. Conclusions The abnormality incidence proportions of blood constituents in university students were estimated to be $0.01-0.08$, and the continual abnormality proportions were estimated to be approximately $0.5-0.8$.


Keywords: university student, blood test, tracking, abnormality incidence proportion, continual abnormality proportion
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## 1. Introduction

As blood tests are not required in regular health examinations for university students, they are seldom performed. However, even in university students, abnormalities such as impaired liver function, serum lipid abnormalities and high serum uric acid are not uncommon, and anemia is a relatively common health problem for young women. In addition, the preliminary conditions of lifestyle diseases have
been shown to begin in childhood [Berenson at al, 1998; Hemmingsson et al, 2006; Komatsu et al, 1996; McMahan et al, 2006; Wamala et al, 2001], and serum lipid abnormalities, etc. of adolescence, etc. are risk factors of arteriosclerosis in adulthood [Daviglus et al, 2004; Kronmal et al, 1993; Stamler et al, 2000]

To examine the value of blood tests in health examinations of university students, it is important to know not only the relationships between blood characteristics and health problems, but also the
dynamic states of blood characteristics in adolescence [Usui et al, 2001; Yasui et al, 2004]. Therefore, using data from $1^{\text {st }}$-year and $4^{\text {th }}$-year blood tests in university students performed over a 3 -year period, we attempted to examine relationships between $1^{\text {st }}$ year and $4^{\text {th }}$-year values and estimate the probability of subjects who are abnormal in one year being abnormal one year later as well as the probability of subjects who are normal in one year becoming abnormal one year later.

## 2. Methods

### 2.1. Definition of terminology

In this paper some terms are defined and abbreviated as follows:
(1) $1^{\text {st }}$-year $-4^{\text {th }}$-year correlation; correlation coefficient between $1^{\text {st }}$-year and $4^{\text {th }}$-year blood test;
(2) $1^{\text {st }}$-year abnormality proportion; abnormality prevalence proportion of blood test at school entry;
(3) $4^{\text {th }}$-year abnormality proportion; abnormality prevalence proportion of blood test at the beginning of the $4^{\text {th }}$ year;
(4)abnormality-abnormality proportion; abnormality prevalence proportion of blood test in the $4^{\text {th }}$ year among subjects with $1^{\text {st }}$-year abnormality;
(5)normality-abnormality proportion; abnormality
prevalence proportion of blood test in the $4^{\text {th }}$ year among subjects with $1^{\text {st }}$-year normality;
(6)continual abnormality proportion; abnormality prevalence proportion of blood test after one year among subjects with abnormality;
(7)abnormality incidence proportion; abnormality prevalence proportion of blood test after one year among subjects with normality.

### 2.2. Correlation Coefficient

We examined relationships between $1^{\text {st }}$ - and $4^{\text {th }}$ year values by simple correlation coefficients. In the calculations, linearity between the values was observed by scatter diagrams.

### 2.3. Analytical model

### 2.3.1. Analytical model 1

The continual abnormality proportion is designated as $\alpha$, the abnormality incidence proportion is designated as $\beta$, and $\alpha$ and $\beta$ are assumed to be the same ratio every year, respectively. If $1^{\text {st }}$-year abnormality proportion is designated as p and $1^{\text {st }}$ year normality proportion is designated as $q$, then in the whole population of subjects who were abnormal at the time of university entrance, the abnormality proportion after 1 year is $\mathrm{p} \alpha$, the normality proportion


Figure 1 Divergence of abnormal subjects at school entry in following years
after 1 year is $p(1-\alpha)$, and the normal and abnormal diverge after 2 and 3 years as shown in Figure 1. Subjects who were normal at the time of university entrance showed the same divergence.

From the above relationship, if the abnormalityabnormality proportion is designated as $r$ and the normality-abnormality proportion is designated $s$, then $\alpha, \beta, r$, and $s$ are expressed by the following relational expressions:

$$
\begin{align*}
& f(\alpha, \beta)=\alpha^{3}+\alpha(1-\alpha) \beta+\alpha(1-\alpha) \beta+(1-\alpha) \beta(1-\beta)=\mathrm{r}  \tag{1}\\
& g(\alpha, \beta)=\alpha^{2} \beta+(1-\alpha) \beta^{2}+\alpha \beta(1-\beta)+\beta(1-\beta)^{2}=\mathrm{s}  \tag{2}\\
& ; 0 \leq \alpha, \beta \leq 1,0 \leq \mathrm{r}, \mathrm{~s} \leq 1
\end{align*}
$$

Since $r$ and $s$ are given as observed values, $\alpha$ and $\beta$ can be calculated by solving the simultaneous equation (1) and (2). Solutions obtained from calculation software Mathematica were plural. However, since there was only one real value satisfying $0 \leq \alpha$ and $\beta \leq 1$ in all of the items, the desired estimated value was obtained for all items.

### 2.3.2. Analytical model 2

Since abnormal and normal subjects come and go every year in analytical Model 1, calculation is somewhat cumbersome. Therefore, we performed estimations by Model 2, which ignored the annual coming and going of abnormal subjects and was, therefore, simple; and we compared the results with those of Model 1.

If $1^{\text {st }}$-year abnormal subjects are assumed to be abnormal the following year at the same ratio every year, then the following relationships hold true:

$$
\begin{align*}
& \mathrm{Ab}_{1} \times \alpha \times \alpha \times \alpha=A b_{14} \\
& \dot{\alpha^{3}}=\mathrm{Ab}_{14} / A b_{1}=\mathrm{r} \\
& \dot{\alpha}={ }^{3} V_{\mathrm{r}} \tag{3}
\end{align*}
$$

; $\mathrm{Ab}_{1}=1^{\text {st }}$-year abnormal subjects,
$\mathrm{Ab}_{14}=4^{\text {th }}$-year abnormal subjects among $1^{\text {stt-year }}$
abnormal subjects
If $1^{\text {st }}$-year normal subjects are assumed to become abnormal at the same ratio every year and continue to be abnormal during the following years, the following
relationships are obtained:

$$
\begin{aligned}
& \mathrm{No}_{1} \times \dot{\beta}+\left(\mathrm{No}_{1}-\mathrm{No}_{1} \times \dot{\beta}\right) \times \dot{\beta}+\left(\mathrm{No}_{1}-\mathrm{No}_{1} \times \dot{\beta}-\left(\mathrm{No}_{1}-\mathrm{No}_{1} \times \dot{\beta}\right) \times\right. \\
& \dot{\beta}) \dot{\beta}=\mathrm{NoAb}_{14} \\
& \mathrm{No}_{1} \times \dot{\beta} \times 3-\mathrm{No}_{1} \times \dot{\beta^{2}} \times 3+\mathrm{No}_{1} \times \dot{\beta^{3}}=\mathrm{NoAb}_{14}
\end{aligned}
$$

Here, since the value of $\beta$ is small, approximately the following relations hold true:
$\mathrm{No}_{1} \times \beta \times 3 \approx \mathrm{NoAb}_{14}$
$\dot{\beta} \approx\left(\mathrm{NoAb}_{14} / \mathrm{No}_{1}\right) / 3=\mathrm{s} / 3$
; $\mathrm{No}_{1}=1^{\text {st }}$-year normal subjects,
NoAb $b_{14}=4^{\text {th }}$-year abnormal subjects among $1^{\text {st }}$ year
normal subjects
$\alpha$ is estimated from formula (3) and $\dot{\beta}$ is estimated from formula (4)

### 2.4. Data

The data used were blood test results from 2461 students ( 1607 males and 854 females) who entered School T of University C from 1995 to 1997.

Eighteen blood test items were analyzed, including total protein (TP), albumin (Alb), albumin/globulin ratio ( $\mathrm{A} / \mathrm{G}$ ratio), aspartate transaminase (AST), alanine aminotransferase (ALT), $\gamma$-glutamyltransferase ( $\gamma$-GTP), urea nitrogen (UN), creatinine (Cr), uric acid (UA), total cholesterol (TC), high-density lipoprotein cholesterol (HDLC), neutral fat (NF), leukocytes (WBC), erythrocytes (RBC), hemoglobin $(\mathrm{Hb})$, hematocrit $(\mathrm{Ht})$, platelets (PLT) and serum iron ( Fe ).

The estimations were performed only for items having 30 or more $1^{\text {st }}$-year abnormal subjects each of males and females.

Blood tests were performed by LIAC, Ltd. Abnormal values were determined using standard values published by the same company, and the results were described as low-TC, high-TC, etc. Data analyses were performed using statistical software SPSS14.0 for Windows.

## 3. Results

### 3.1. Correlations between $1^{\text {st }}$-Year and $4^{\text {th }}-$ Year Values

The $1^{\text {st }}-y e a r-4^{\text {th }}$-year correlations are shown in
Table 1. The correlation coefficients of all items showed significant differences at $\mathrm{p}<0.01$. Items with high correlations were UA (males), TC (males) and PLT (males and females) $(0.7 \leq)$; and $\gamma$-GTP (males), Cr (females), TC (females) and RBC (males and females) ( $0.6 \leq,<0.7$ ). Items with low correlations were AST (males), ALT (males and females) and Fe (males and females) ( $<0.3$ ). Items with large differences between males and females were AST, $\gamma$-GTP and UA.

## 3.2. $1^{\text {st }}$-Year and $4^{\text {th }}$-Year Abnormality Proportions

The $1^{\text {stt}}$-year and $4^{\text {th }}$-year abnormality proportions of blood tests are shown in Table 2. The $1^{\text {st }}$-year abnormality proportions were 0.02 or higher for all items. Items showing the proportions of 0.1 or higher were low-TC (male and female), low-HDLC (males), high-NF (males), low-NF (females) and low-Fe (females).

The $4^{\text {th }}$-year abnormality proportions were 0.02
or higher for all items. High-A/G (males), high-TC (females), low-TC (males), low-HDLC (males), highNF (males) and low-Fe (females) had 0.1 or higher proportions.

The $4^{\text {th }}$-year abnormality proportions showed higher values than the $1^{\text {st }}$-year abnormality proportions for two-thirds of all items. Items that revealed a large difference between $1^{\text {sts}}$-year and $4^{\text {th }}$-year proportions were high-TP ( $1^{\text {st }}$ year $>4^{\text {th }}$ year), high-A/G ratio ( $4^{\text {th }}$ year $>1^{\text {st }}$ year) and low-NF (female) $\left(1^{\text {st }}\right.$ year $>4^{\text {th }}$ year).

### 3.3. The $4^{\text {th }}$-Year Abnormality Proportions According to the Status in the 1st Year.

Table 3 shows the abnormality-abnormality proportions (r) and normality-abnormality proportions (s).

The highest abnormality-abnormality proportion (r) was high-TC (males) at 0.5526 , while the lowest was low-NF (males) at 0.0748 . Items that had the proportions of 0.4 or more were high-ALT (males), high-TC, low-TC (males and females), low-HDLC (males) and high-NF (males).

Among the normality-abnormality proportions (s), high-NF (males) was the highest, followed by low-Fe (females), high-A/G (males), and low-HDLC (males),

Table 1 Correlation coefficients between $1^{\text {st }}$-year values and $4^{\text {th }}$-year values

| Item | Males ( $\mathrm{N}=1607$ ) | Females ( $\mathrm{N}=854$ ) |
| :---: | :---: | :---: |
| TP | $0.48{ }^{* *}$ | $0.48{ }^{* *}$ |
| Alb | 0.40 ** | 0.40 ** |
| A/G ratio | 0.43 ** | 0.50 ** |
| AST | 0.10 ** | 0.36 ** |
| ALT | $0.22{ }^{* *}$ | 0.19 ** |
| $\gamma$-GTP | 0.65 ** | 0.38 ** |
| UN | $0.46{ }^{* *}$ | $0.42{ }^{* *}$ |
| Cr | $0.55{ }^{* *}$ | 0.63 ** |
| UA | 0.71 ** | 0.59 ** |
| TC | $0.74{ }^{* *}$ | 0.62 ** |
| HDLC | 0.57 ** | 0.55 ** |
| NF | $0.36{ }^{* *}$ | $0.35{ }^{* *}$ |
| WBC | 0.40 ** | 0.45 ** |
| RBC | $0.67{ }^{* *}$ | $0.66{ }^{* *}$ |
| Hb | 0.57 ** | 0.57 ** |
| Ht | 0.53 ** | 0.54 ** |
| PLT | $0.73{ }^{* *}$ | $0.74{ }^{* *}$ |
| Fe | $0.14{ }^{* *}$ | 0.17 ** |

[^0]Table $21^{\text {st }}$-year abnormality proportions and $4^{\text {th }}$-year abnormality proportions

| Item• Abnormality $\cdot$ Sex |  |  | $1^{\text {st }} \text {-year }$ <br> Abnormality | $4^{\text {th }}$-year <br> Abnormality | Risk Ratio |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Proportion | Proportion |  |
| TP | High | M | 0.0784 (126) | 0.0386 (62) | 0.4923 |
|  |  | F | 0.0761 (65) | 0.0363 (31) | 0.4770 |
| A/G Ratio | High | M | 0.0205 (33) | 0.1400 (225) | 6.8293 |
| AST | High | M | 0.0268 (43) | 0.0243 (39) | 0.9067 |
| ALT | High | M | 0.0423 (68) | 0.0660 (106) | 1.5603 |
| UN | Low | F | 0.0386 (33) | 0.0667 (57) | 1.7280 |
| UA | High | M | 0.0255 (41) | 0.0299 (48) | 1.1725 |
| TC | High | M | 0.0473 (76) | 0.0635 (102) | 1.3425 |
|  |  | F | 0.0843 (72) | 0.1042 (89) | 1.2361 |
|  | Low | M | 0.2757 (443) | 0.2016 (324) | 0.7312 |
|  |  | F | 0.1183 (101) | 0.1101 (94) | 0.9307 |
| HDLC | High | F | 0.0562 (48) | 0.0597 (51) | 1.0623 |
|  | Low | M | 0.1220 (196) | 0.1406 (226) | 1.1525 |
| NF | High | M | 0.1070 (172) | 0.1792 (288) | 1.6748 |
|  |  | F | 0.0480 (41) | 0.0679 (58) | 1.4146 |
|  | Low | M | 0.0666 (107) | 0.0212 (34) | 0.3183 |
|  |  | F | 0.1206 (103) | 0.0429 (69) | 0.3557 |
| WBC | High | M | 0.0243 (39) | 0.0249 (40) | 1.0247 |
|  | Low | M | 0.0299 (48) | 0.0398 (64) | 1.3311 |
| RBC | High | M | 0.0454 (73) | 0.0361 (58) | 0.7952 |
| Fe | High | M | 0.0672 (108) | 0.0790 (127) | 1.1756 |
|  | Low | M | 0.0716 (115) | 0.0784 (126) | 1.0950 |
|  |  | F | 0.1557 (133) | 0.1838 (157) | 1.1805 |

Note) The estimations were performed for items having 30 or more abnormal subjects in the $1^{\text {st }}$ year for males and females, respectively.
Number in parentheses: Number of abnormal subjects
Risk Ratio: $4^{\text {th }}$-year Abnormality Proportion $/ 1^{\text {st }}$-year Abnormality Proportion
each at values of 0.1 or higher. The lowest proportion was low-NF (males).

For all items, the abnormality-abnormality proportions(r) were higher than the normalityabnormality proportions (s). Risk ratios (r/s) were particularly high for high-RBC (males), high-UA (males), high-TC (males), high-ALT (males), etc. The ratios were low for Fe .

### 3.4. Estimation of Continual Abnormality Proportions ( ${ }^{\mathfrak{\beta}} \downarrow \mathbf{r}, \boldsymbol{\alpha}$ ) and Abnormality Incidence Proportions ( $\mathbf{s} / \mathbf{3}, \boldsymbol{\beta}$ )

The results of estimation are shown in Table 4. Except for low-NF (males), high-Fe (males) and lowFe (males) by Model $1(\alpha)$, and except for low-NF (males) by Model $2\left({ }^{3} \sqrt{ }\right.$ r), all continual abnormality proportions were 0.5 or higher. Items that showed
continual abnormality proportions of 0.7 or more by Model 1 were high-ALT (males), low-TC, high-TC, low-HDLC (males), high-NF (males) and high-RBC (males). Items that showed continual abnormality proportions of 0.7 or more by Model 2 were high-A/ G ratio (males), high-ALT (males), high-TC, lowTC, low-HDLC (males), high-NF (males), highRBC (males) and low-Fe (females). Continual abnormality proportions of Fe were lower than other items; however, the low values of females were approximately 0.7 . Estimated values were higher by Model 2 than by Model 1 for all items. However, except for low-Fe (males), ratios of estimates $\left({ }^{3} V_{\mathrm{r}} / \alpha\right)$ were less than 1.1.

Abnormality incidence proportions were less than 0.1 for both Model $1(\beta)$ and Model $2(\mathrm{~s} / 3)$. The estimates, unlike those of the continual abnormality proportions, were greater by Model 1 than by Model

Table 3 Abnormality proportions after 3 years according to the $1^{\text {st }}$-year states

| Item• Abnormality $\cdot$ Sex |  |  | $1^{\text {st }}$-Year States |  | r/s |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Abnormal( r ) | Normal( s ) |  |
| TP | High | M | 0.2143(27/126) | 0.0236(35/1480) | 9.0805 |
|  |  | F | 0.1846 (12/65) | 0.0241 (19/788) | 7.6598 |
| A/G Ratio | High | M | 0.3636 (12/33) | 0.1348(212/1573) | 2.6973 |
| AST | High | M | 0.1628 (7/43) | 0.0205(32/1560) | 7.9415 |
| ALT | High | M | 0.5000 (34/68) | 0.0462(69/1493) | 10.8225 |
| UN | Low | F | 0.2727 (9/33) | 0.0579 (47/812) | 4.7098 |
| UA | High | M | 0.2683 (11/41) | 0.0239(37/1547) | 11.2259 |
| TC | High | M | 0.5526 (42/76) | 0.0551(60/1088) | 10.0290 |
|  | Low | F | 0.4861 (35/72) | 0.0778 (53/681) | 6.2481 |
|  |  | M | 0.5102(226/443) | 0.0901(98/1088) | 5.6626 |
|  |  | F | 0.4752(48/101) | 0.0675 (46/681) | 7.04 |
| HDLC | High | F | 0.3125 (15/48) | 0.0457 (36/787) | 6.8381 |
|  | Low | M | 0.4133(81/196) | $0.1049(145 / 1382)$ | 3.9399 |
| NF | High | M | 0.4070(70/172) | 0.1581(210/1328) | 2.5743 |
|  | Low | F | 0.2439 (10/41) | 0.0676 (48/710) | 3.6080 |
|  |  | M | 0.0748 (8/107) | 0.0181(24/1328) | 4.1326 |
|  |  | F | 0.2039(21/103) | 0.0662 (47/710) | 3.0801 |
| WBC | High | M | 0.1538 (6/39) | $0.0224(34 / 1520)$ | 6.8661 |
|  | Low | M | 0.2500 (12/48) | 0.0342 (52/1520 | 7.3099 |
| RBC | High | M | 0.3562 (26/73) | $0.0209(32 / 1531)$ | 17.0431 |
| Fe | High | M | $0.1481(16 / 108)$ | 0.0744(103/1384) | 1.9906 |
|  | Low | M | 0.1652(19/115) | 0.0715(99/1384) | 2.3105 |
|  |  | F | 0.3459(46/133) | $0.1546(107 / 692)$ | 2.2374 |

Note) Abnormal(r): Abnormality-abnormality proportion Normal(s): Normality-abnormality proportion

2 for all items, and the differences between the values estimated by both models were greater for the abnormality incidence proportions than for the continual abnormality proportions. Items that showed comparatively high values by Model 1 were high-A/

G ratio (males), high-NF (males), low-Fe (females), etc. Items that showed small differences between Model 1 and Model 2 were TC, HDLC, etc., and an item with large differences between the models was Fe. However, even for high-TC (males) that showed a

Table 4 Estimates of continual abnormality proportions $\left(\alpha,{ }^{3} \sqrt{ } \mathrm{r}\right)$ and abnormality incidence proportions $(\beta, \mathrm{s} / 3)$

| Item•Abnormality•Sex |  |  | $\alpha$ | $\beta$ | ${ }^{3} \sqrt{ } \mathrm{r}$ | $\mathrm{s} / 3$ | ${ }^{3} V_{\mathrm{r}} / \alpha$ | s/3/ $\beta$ | $\alpha / \beta$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| TP | High | M | 0.5880 | 0.0124 | 0.5984 | 0.0079 | 1.0177 | 0.6371 | 47.42 |
|  |  | F | 0.5566 | 0.0131 | 0.5694 | 0.0080 | 1.0230 | 0.6107 | 42.49 |
| A/G Ratio | High | M | 0.6795 | 0.0670 | 0.7137 | 0.0449 | 1.0503 | 0.6701 | 10.14 |
| AST | High | M | 0.5335 | 0.0114 | 0.5460 | 0.0068 | 1.0234 | 0.5965 | 46.80 |
| ALT | High | M | 0.7881 | 0.0196 | 0.7937 | 0.0154 | 1.0071 | 0.7857 | 40.21 |
| UN | Low | F | 0.6285 | 0.0296 | 0.6485 | 0.0193 | 1.0318 | 0.6520 | 21.23 |
| UA | High | M | 0.6371 | 0.0119 | 0.6450 | 0.0080 | 1.0124 | 0.6723 | 53.54 |
| TC | High | M | 0.8151 | 0.0228 | 0.8206 | 0.0184 | 1.0067 | 0.8070 | 35.75 |
|  | Low | F | 0.7758 | 0.0339 | 0.7863 | 0.0259 | 1.0135 | 0.7640 | 22.88 |
|  |  | M | 0.7880 | 0.0390 | 0.7991 | 0.0300 | 1.0141 | 0.7692 | 20.21 |
|  |  | F | 0.7710 | 0.0295 | 0.7804 | 0.0225 | 1.0122 | 0.7627 | 26.14 |
| HDLC | High | F | 0.6660 | 0.0222 | 0.6786 | 0.0152 | 1.0189 | 0.6847 | 30.00 |
|  | Low | M | 0.7248 | 0.0492 | 0.7449 | 0.0345 | 1.0277 | 0.7012 | 14.73 |
| NF | High | M | 0.7071 | 0.0781 | 0.7411 | 0.0527 | 1.0481 | 0.6748 | 9.05 |
|  | Low | F | 0.5968 | 0.0361 | 0.6248 | 0.0225 | 1.0469 | 0.6233 | 16.53 |
|  |  | M | 0.3960 | 0.0118 | 0.4213 | 0.0060 | 1.0639 | 0.5085 | 33.56 |
|  |  | F | 0.5535 | 0.0371 | 0.5886 | 0.0221 | 1.0634 | 0.5957 | 14.92 |
| WBC | High | M | 0.5211 | 0.012 | 0.5358 | 0.0075 | 1.0282 | 0.5906 | 41.03 |
|  | Low | M | 0.6173 | 0.0175 | 0.6300 | 0.0114 | 1.0206 | 0.6514 | 35.27 |
| RBC | High | M | 0.7043 | 0.0096 | 0.7089 | 0.0070 | 1.0065 | 0.7292 | 73.36 |
| Fe | High | M | 0.4659 | 0.0466 | 0.5291 | 0.0248 | 1.1357 | 0.5322 | 10.00 |
|  | Low | M | 0.4973 | 0.0431 | 0.5487 | 0.0238 | 1.1034 | 0.5522 | 11.54 |
|  |  | F | 0.6572 | 0.0810 | 0.7020 | 0.0515 | 1.0682 | 0.6358 | 8.11 |

small difference, the ratio of estimate $(\mathrm{s} / 3 / \beta)$ was 0.8 , which indicates that the estimate by Model2 is less than that by Modell by approximately 20\%. HighAST, low-NF, high-WBC (males), Fe, etc. showed large differences between the estimates of two models with approximately $0.5-0.6$ of the $s / 3 / \beta$.

## 4. Discussion

Using the blood tests performed over a 3-year period, we attempted to analyze the variability of blood constituents by correlation coefficients, and estimates of continual abnormality proportions
and abnormality incidence proportions. Items with comparatively high correlations were $\mathrm{Cr}, \mathrm{UA}, \mathrm{TC}$, HDLC, RBC, $\mathrm{Hb}, \mathrm{Ht}$ and PLT, and items with low correlations were AST, ALT and Fe. For continual abnormality proportions, no significant differences were found between the estimates by Model 1, which assumed annual coming and going of normal and abnormal subjects, and by Model 2, which ignored coming and going. For abnormality incidence proportions, however, Model 2 tended to yield smaller values than Model 1.

The reason that large differences were not seen in the continual abnormality proportions of the two models is thought to be that the abnormality incidence proportions of the blood test items covered here were much lower than the continual abnormality proportions. Conversely, the large differences observed in the abnormality incidence proportions between the models were probably caused by the fact that the abnormality incidence proportions were low and were, therefore, more easily influenced by abnormal subjects who became normal the next year and abnormal the next year. For many blood test items, the abnormality incidence proportion is generally lower than the continual abnormality proportion. Therefore, models that assume the coming and going of abnormal subjects and normal subjects are better for estimation of abnormality incidence proportion. On the other hand, it is thought that the continual abnormality proportion is not influenced much by coming and going due to low abnormality incidence proportion, and that relatively good estimates can also, therefore, be obtained by the simple model that does not assume the coming and going.

Model 1 assumes that the probability that normal subjects will become abnormal after 1 year and the probability that abnormal subjects will be abnormal after 1 year are constant every year, respectively; however, continually normal subjects and normal subjects who were previously abnormal are naturally included among the normal subjects. Continually normal subjects are generally thought to have a lower probability of becoming abnormal. In addition, continually abnormal subjects and subjects who became abnormal that year are also included among abnormal subjects. In this case as well, continually abnormal subjects probably have a higher probability of becoming abnormal. Therefore, it is necessary to emphasize that the estimates of this report are based
on these assumptions.
Abnormality incidence proportions based on Model 1 were approximately $0.01-0.08$, and continual abnormality proportions were 0.8 for high items and 0.5-0.7 for many items. Reports on the dynamic states of blood constituents in young adulthood are relatively few, and no reports were found on abnormality incidence proportion. Even for issues of continual abnormality proportion, although there are quite a few reports on tracking phenomena [Boulton et al, 1995;Fuentes et al, 2003; Kelder et al, 2002; Moilanen et al, 1987; Omura et al, 1991; Patel et al, 2007; Porkka et al, 1994; Spyckerelle et al, 1992; Webber et al, 1991], no reports were found that addressed continual abnormalities after 1 year. Table 5 summarizes the reports on tracking phenomena [Boulton et al, 1995; Fuentes et al, 2003; Kelder et al, 2002; Moilanen et al, 1987; Omura et al, 1991; Porkka et al, 1994; Spyckerelle et al, 1992; Webber et al, 1991].

According to these reports, serum lipids during childhood have correlation coefficients between 0.45 0.55 (5 years) and 0.71-0.78 (2 years). Relatively high correlations are observed even when the observation interval is long. The correlation coefficients of NF in this report are somewhat low; however, the correlation coefficients of TC and HDLC closely resemble the values of these reports. NF fluctuates to a relatively high degree because of the influence of meals, etc. Meals before blood sampling were not restricted this time; therefore, this might have influenced the low correlation of NF. No reports were found on the tracking phenomena of blood constituents other than serum lipids; consequently, adequate bibliographical consideration is not possible. However, except for the fact that TP, WBC, Ht, etc. are affected greatly by exercise, there is no mention in books, etc. of the items covered here fluctuating greatly during everyday life. Therefore, we think that the correlations of $1^{\text {stt}}$-year and $4^{\text {th }}$-year values present a picture of tracking of blood constituents to some extent.

Except for $\mathrm{A} / \mathrm{G}$ ratio and NF, the $1^{\text {st }}$-year and $4^{\text {th }}$-year abnormality proportions do not differ by more than two times. This fact does not mean the turnover between normal subjects and abnormal subjects is low. Actually, the abnormality-abnormality proportions are approximately 0.55 for high items and approximately 0.1 for low items, and at least half of these rates are 0.3 or less. Furthermore, the estimates

Table 5 Summary of literature on tracking phenomena of blood constituents and other characteristics in young subjects

| Authors | Subjects | Term | Item | Correlation Coefficient |
| :---: | :---: | :---: | :---: | :---: |
| Fuentes et al | 7-15 years old | 8 years interval | Total cholesterol | 0.655 |
| Kelder et al | 3-8 school year | 6 years interval | Body weight <br> BMI <br> Skinfold thickness <br> Serum lipids <br> Blood pressure | 0.86 0.86 $0.72-0.78$ $0.67-0.72$ $0.45-0.51$ |
| Boultonet al | 13-15 years old | 2 years interval | Serum lipids | 0.71-0.78 |
| Omura et al | Junior high school Student (First time test) | 5 years interval | Total cholesterol <br> Systolic blood pressure <br> Diastolic blood pressure | $\begin{aligned} & 0.45-0.55 \\ & 0.28-0.33 \\ & 0.19-0.36 \end{aligned}$ |
| Webber et al | Childhood - <br> Adulthood | 12 years interval | Serum lipids | $\begin{gathered} \text { LDLC }> \\ \text { TC,HDLC } \end{gathered}$ |
| Moilanen et al | 3-18 years old(First time test) | 3 years interval | Serum linoleic acid arachidonic acid | $\begin{aligned} & 0.58 \\ & 0.61 \end{aligned}$ |
| Spyckerelle et al | 4-17 years old(First time test) | 5 years interval | Total cholesterol Uric acid Glucose | Over 0.5 <br> Over 0.5 <br> About 0.3 |
| Porkka et al | 9 years old(First time <br> test ) <br> Male | 3 years <br> 6 years <br> 9 years | Total cholesterol | $\begin{aligned} & 0.70 \\ & 0.73 \\ & 0.59 \end{aligned}$ |

for continual abnormality proportions were 0.8 for high items and 0.4 for low items, and at least half of these rates are 0.7 or less. These results indicate that for many items, $30 \%$ or more of abnormal subjects become normal after 1 year and $70 \%$ or more of abnormal subjects after 3 years.

The turnover between abnormal and normal subjects occurs considerably often; however, there is a great difference between the probability of abnormal subjects remaining abnormal (continual abnormality proportion) and of normal subjects becoming abnormal (abnormality incidence proportion). The abnormality-abnormality proportions were greater than 5 times higher than the normality-abnormality proportions for at least half of the items. Comparing the continual abnormality proportions ( $\alpha$ ) with abnormality incidence proportions ( $\beta$ ), the difference is even greater. $\alpha / \beta$ value was low for Fe at about 10 , and high for high-RBC at 73 , and 20 or more for many items. These results show that even in young
subjects, the probability of continual abnormalities is much higher than that of new abnormalities.

In light of the estimated values of continual abnormality proportion and abnormality incidence proportions, the dynamic state of blood characteristics for young subjects is thought to differ considerably by item. Serum lipids have high continual abnormality proportions and also have high abnormality incidence proportions. During the period from puberty until young adulthood, lipid metabolism changes to relatively high degree [Berenson et al, 1981; Porkka et al, 2008], and this change may be manifested in a relatively high abnormality incidence proportion. On the other hand, continual abnormality proportions are low and abnormality incidence proportions are high for Fe . The $1^{\text {st}}$-year- $4^{\text {th }}$-year correlations for Fe are also low, and therefore, serum iron is likely to be characterized by large fluctuations. Incidentally, RBC, which has a close relation to Fe , showed high continual abnormality proportions, low abnormality
incidence proportions and correlation coefficients near 0.7 . The dynamic states of Fe and RBC are thought to differ considerably.

The data we used were the results of blood tests carried out for all students entering School T of University C from 1995 to 1997. Among the subjects determined to be abnormal during the $1^{\text {st }}$ year, subjects with severe abnormalities received second examinations. Some of these subjects, although few, were treated for anemia and impaired liver function. Therefore, it is difficult to rule out the possibility that the results relating to $\mathrm{RBC}, \mathrm{Fe}, \mathrm{AST}$ and ALT were influenced by treatment.

## 5. Conclusions

Approximately the same continual abnormality proportions were obtained from both models. Even the simple model that does not assume the coming and going of abnormal subjects provides good estimates. Blood constituents of university students had the following characteristics. Items with high 1st-year-4th-year correlations were TC, etc., and items with low correlations were AST, ALT, and Fe. The 4th-year abnormality proportions were higher than the 1 st-year abnormality proportions for many items; however, with the exception of serum protein and NF, the differences were not significant. The 4thyear abnormality proportions of 1st-year abnormal subjects, compared to the 4th-year abnormality proportions of 1st-year normal subjects, were twice as high for low items and over 10 times as high for high items. The continual abnormality proportions of blood constituents in university students were estimated to be approximately $0.5-0.8$, and the abnormality incidence proportions were estimated to be 0.01-0.08

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[^0]:    **: $\mathrm{P}<0.01$

